

# A Costly Commitment: Populism, Economic Performance, and the Quality of Bureaucracy\*

Luca Bellodi<sup>†</sup>

Massimo Morelli<sup>‡</sup>

Matia Vannoni<sup>§</sup>

## Abstract

We study the consequences of populism for economic performance and the quality of bureaucracy. When voters lose trust in representative democracy, populists strategically supply unconditional policy commitments that are easier to monitor for voters. When in power, populists try to implement their policy commitments regardless of financial constraints and expert assessment of the feasibility of their policies, worsening government economic performance and dismantling resistance from expert bureaucrats. With novel data on more than 8,000 Italian municipalities covering more than 20 years, we estimate the effect of electing a populist mayor with a close-election regression discontinuity design. We find that the election of a populist mayor leads to more debts, a larger share of procurement contracts with cost overruns, higher turnover among top bureaucrats – driven by forced rather than voluntary departures – and a sharp decrease in the percentage of graduate bureaucrats.

**Keywords:** Populism, Economic Performance, Bureaucracy, Regression Discontinuity Design.

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<sup>†</sup>Bocconi University and DONDENA Center.

<sup>‡</sup>Bocconi University, IGIER, Baffi CAREFIN Center, and CEPR.

<sup>§</sup>King’s College London and Baffi CAREFIN Center.

*“After stopping the landings, my greatest commitment now is for the expulsion of those who do not flee from any war, but who bring war to our home, who shout insults towards the country that is hosting them, and demand, demand, demand!  
I don’t give up, I assure you!  
#stopinvasion”*

Mr. Matteo Salvini, December 15, 2018

This is one of the many tweets that Matteo Salvini, leader of the League, posted during his fight against immigration as Minister of the Interior for the Italian Government between June 2018 and September 2019.<sup>1</sup> Other populist leaders, both in Italy and around the world, centred their political campaigns on similar promises: 5-Star-Movement founder Beppe Grillo was the strongest advocate for a “citizenship income” to protect those left out of the labour force as a result of automation and economic growth,<sup>2</sup> President Jair Bolsonaro repeatedly galvanised gun ownership as a mean to fight crime and violence in Brazil,<sup>3</sup> and the symbol of Donald Trump’s anti-immigration policy was the wall at the US-Mexico border, which he promised during his inaugural speech as presidential candidate: *“I will build a great great wall on our southern border and I’ll have Mexico pay for that wall”*.<sup>4</sup> These statements shed light on a distinctive yet unexplored feature that characterises both populist campaigns and policies: unconditional commitments. In this paper we show how populists’ commitments are costly, with detrimental consequences for economic performance and bureaucratic quality.

A growing literature studies the cultural, economic, technological, and political *causes* of populism (for recent reviews, see [Guriev and Papaioannou 2022](#); [Noury and Roland 2020](#); [Berman 2021](#)), but less is known about the *consequences* of populism for policy-making. There are two challenges researchers are confronted with when studying the effects of populism: one theoretical, one empirical.

The theoretical challenge consists of deriving expectations from a logical argument on why populism should be good or bad for the economy and government performance. [Dornbusch and Edwards \(1991\)](#), in their study of populism in Latin America, make a clear association between populism and specific types of expansionary and redistributive policies and argue that,

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<sup>1</sup>Tweet in Italian at <https://twitter.com/matteosalvinimi/status/1074036205052968960>.

<sup>2</sup>See for instance <https://beppegrillo.it/basic-income-now-is-the-time/>.

<sup>3</sup>See <https://www.nytimes.com/2020/03/31/world/americas/guns-brazil-bolsonaro.html>.

<sup>4</sup>Full text available at <https://www.washingtonpost.com/news/post-politics/wp/2015/06/16/full-text-donald-trump-announces-a-presidential-bid/>.

by neglecting constraints and inflationary consequences, populism leads to economic failure, social unrest, and violence. In their words, because of specific policy preferences, populism is “self destructive.” However, recent scholarship questioned the association between populism and specific types of policies, characterising populism as a thin ideology with a clear view of society but a vague view of the ideal economic or political system (Mudde and Kaltwasser 2017). When departing from a fixed association with policies, the mechanism behind the negative consequences of populism escapes and the idea that populists are doomed to bad performance seems undistinguishable from a mere prejudice. In fact, it has been shown that in some cases populists can be successful at governing (Albertazzi and McDonnell 2015; Biard 2019). The second challenge is empirical. Estimating the effect of populism is a daunting task. Random occurrences of populist governments are nearly inexistent, and multiple economic and political factors that are correlated with the election of a populist leader also correlate with policy outcomes, limiting the space for causal identification. With this paper we address both challenges.

Theoretically, we build on recent work that looks at the determinants of populists’ incentives and policy platforms. When economic, technological, or cultural shocks erode citizens’ trust in political elites and the traditional tools of representative democracy, voters demand simple protection policies (see e.g. Guiso et al. 2017). Populist politicians intercept this demand shift towards tighter control over policy-making, and run for office *i*) committing to a set of policies easy to monitor (e.g., America First, or citizenship income), and *ii*) fuelling voters’ distrust in the non-populist opponents, accused of being supported by corrupted elites (Moffitt 2016; Morelli, Nicolò, and Roberti 2021). Trust erosion therefore triggers populism, for it creates incentives for politicians to commit to policies that are easier to monitor, while at the same time accusing traditional parties’ candidates of being captured by the elites. Because of such commitments, when populists come to power they implement their agenda with no regard for constraints and expert assessment of the feasibility and consequences of their policies. When facing the resistance of expert bureaucrats’, populists will replace them with non-experts in order to ensure the smooth implementation of their policy agenda (Eichengreen 2018; Sasso and Morelli 2021).

Building on these micro foundations of populists’ incentives and behaviour, we derive

testable expectations about the consequences of populism for policy-making. No matter what the circumstances require, populists carry on with their policy commitments, with detrimental consequences for government economic performance. In order to minimise sabotage from a competent bureaucracy, populists replace experts with non-expert bureaucrats, which results in increased turnover and lower quality of bureaucracy.

Empirically, we test these expectations with novel data on economic performance, public procurement, and bureaucratic composition of municipal governments in Italy over a 20-year period, from 1998 to 2020. We use a close-election regression discontinuity design and compare municipalities in which a populist mayor barely won the elections to municipalities where a populist barely lost in order to isolate the effect of electing a populist candidate. Italian local government is a good case to study the effects of electing a populist government. First, Italy has more than 8,000 municipalities and several populist parties populate the Italian party system at both national and local level, yielding large sample size and large variation in “treatment status.” Second, populist parties exert a tight discipline over their members, hence there is likely to be a precise match between the populist nature of the party and the populist attribute of affiliated mayors. Third, the concentration of executive power in mayors and mayors’ discretion over the appointments and removals and top bureaucrats create the necessary space to detect sizeable effects on economic performance and the quality of bureaucracy.

The results support the expectations. We find that the (close) election of a populist mayor leads to larger debt accumulation and smaller debt repayment, as well as to a 4.8 percentage point increase in the share of public contracts with cost overruns. Turnover among top bureaucrats increases by 50% compared to the average turnover in the data, and the percentage of bureaucrats with a university degree drops by -13.1 percentage points. By analysing the stated reasons of bureaucrats’ departures, we are also able to show that forced rather than voluntary departures drive the increased turnover, suggesting that bureaucrats are forced to leave and do not choose to leave when populists win.

Our findings complement early attempts at investigating the relationship between populism, the economy, and bureaucracy. [Funke, Schularick, and Trebesch \(2020\)](#) assemble a dataset on populist leaders and find large long-term economic costs of populism, with GDP per capita being more than 10% below compared to the most plausible non-populist counterfactual. Sim-

ilarly, [Magud and Spilimbergo \(2021\)](#) find negative macroeconomic consequences of populist regimes in Latin American countries. On the bureaucracy side, [Peters and Pierre \(2020\)](#) suggest that populism is likely to translate into lower expertise in government. [Bauer and Becker \(2020\)](#) discuss the public administration goals and strategies of populist governments, showing how purges of personnel and top bureaucrats occurred in many historical cases of populist governments. Similar conclusions are reached in the theory proposed by [Sasso and Morelli \(2021\)](#). Finally, our work builds on recent scholarship showing that populists can have a significant influence on policy, both when in power ([Albertazzi and McDonnell 2015](#); [Paxton 2019](#)), and when outside government ([Biard, Bernhard, and Beta 2019](#)). We push those findings ahead showing that the influence of populist commitments has negative consequences for the performance of municipalities and the retention of competent bureaucrats.

## **Populists' Political Agency and the Strategic Supply of Commitments**

Social scientists have embraced different definitions of populism, either focusing on ideology, political strategy, or policies ([Mudde and Rovira Kaltwasser 2018](#)). First, the “ideational” approach defines populism as a thin-centred ideology which portrays society as divided into the “pure people” and the “corrupt elite,” arguing that politics should be the expression of the will of the people ([Mudde 2004](#)). Second, the political-strategic approach considers populism as a political strategy employed by politicians in the attempt at “winning and exercising power” ([Weyland 2001, 12](#)) while creating unmediated ties between the leader and voters ([Weyland 2017](#)). Consistently with these two definitions, the people-vs-elite rhetoric is a straightforward manifestation of populism, which is also the most common measure used in empirical work (e.g., [Pauwels 2011](#); [Hawkins et al. 2019](#); [Gennaro, Lecce, and Morelli 2021](#); [Di Cocco and Monechi 2021](#)). A third approach, originating from economics, defines populism as the implementation of redistributive policies that claim to promote the interest of common citizens without regard to the consequences of such policies ([Dornbusch and Edwards 1991](#); [Acemoglu, Egorov, and Sonin 2013](#)).

Our theoretical argument endorses a strategic definition of populism, which we believe is the most suitable for studying the consequences of populism. In fact, while the ideational ap-

proach does not point to any specific direction when studying the consequences of populism, the economic approach is too narrow in associating populism with a set of redistributive policies (Kaltwasser et al. 2017). Conversely, the strategic approach allows us to link causes, strategies, and consequences of populism in a single account. As we propose in this section, trust erosion creates demand for simple commitments that are easier to monitor for voters (*causes*) and triggers a strategic response from politicians: unconditional commitments to policies and anti-elite rhetoric (*strategies*). Because unconditional commitments are by definition unfit for changing circumstances and hostile to expert assessment of alternatives, populism has detrimental effects on policy-making (*consequences*).

A common thread in the literature on the causes of populism is the erosion of voters' trust in the traditional tools of representative democracy (Inglehart and Norris 2019; Berman 2021). When trust is high, politicians are entrusted to adjust policy-making to changing circumstances, in line with a *trustee model* of representation (Fox and Shotts 2009). Political selection prioritises competence and when running for office, politicians will stress their level of expertise in delivering their proposed policies. Conversely, when trust is eroded, voters consider politicians – even the competent ones – undeserving of the autonomy characterising the trustee model of representation, perhaps because they are considered selfish rent-seekers captured by interest groups or elites. Voters want to take back control over policy, demanding simple policy commitments that are easy to monitor. Representation shifts to a *committed delegate model*, where voters delegate discretion to politicians insofar as it is compatible with the limited trust. Populist politicians intercept this demand and run for office with simple commitments and radical solutions (e.g., build-wall and zero-immigration types of policies, protectionism, citizenship income). Once a party or politician shifts to the committed delegate model, the best complementary strategy to maximise electoral support is to fuel distrust for the corrupted candidates supported by the corrupted elites, thus sustaining the crisis of trust (see e.g., Mudde 2004; Moffitt 2016). This is why the committed delegate model triggers populism, for the committed delegate rationally chooses simple policy commitments and all the complementary strategies commonly associated with populist behaviour, namely anti-elite rhetoric, anti-media, and anti-experts denigration (Morelli, Nicolò, and Roberti 2021).

Once in office, populists try to stick to their policy commitments. Paxton (2019), for

instance, shows that even at local government-level populists are highly incisive with respect to the policy issues they prioritise, especially when they enjoy a large degree of discretion, as evidenced by the author’s comparative analysis of municipal governments in Italy and Austria. Also cross-country evidence points to the significant influence that populist parties have on policies (Biard, Bernhard, and Beta 2019) and how they manage to deliver on their electoral promises, ultimately surviving the electoral cost of governing (Albertazzi and McDonnell 2015). However, unconditional commitments might lead to inefficiencies and suboptimal outcomes if populist policies do not match the state of the world and are implemented without a careful assessment of their correspondence to the socio-economic landscape. As a result of this form of imperative policy-making, we expect populism to worsen government economic performance.

Furthermore, because commitment-type policies are by definition incompatible with adjustments or alternatives, populist politicians are not willing to have their proposed policies undergo expert assessments of their feasibility, for experts are more likely to raise concerns about the desirability of populists’ policies given the economic and political context (Peters and Pierre 2020). This might be one of the reasons why populist governments were slower at passing health measures and mobility restrictions during the outbreak of the COVID-19 pandemic in February and March 2020 (Kavakli 2020), and in general performed more poorly (Bayerlein et al. 2021). For populists, expertise is less crucial, and expert bureaucrats’ recommendations implying policy solutions different from or not included in their commitment-type agenda are unwelcome (Mudde and Kaltwasser 2017; Eichengreen 2018). Populists therefore have a strict preference for non-expert bureaucrats who will obsequiously implement the populist agenda. As a result, populists will control the bureaucracy by staffing the administration with non-experts who are more likely to implement populist policies without hampering populists’ agenda (Sasso and Morelli 2021).<sup>5</sup>

Based on this theoretical framework, we can derive testable expectations about the consequences of populists’ commitments. As a result of this form of imperative policy-making, we expect populism to worsen government economic performance. Furthermore, when facing resistance from expert bureaucrats, populists will replace them with non-experts, increasing

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<sup>5</sup>In some cases, experts might also “feign” to be non-experts while waiting out the incumbent government (Sasso and Morelli 2021; Cameron and Figueiredo 2020). Because feigning behaviours are hardly tractable empirically, we focus on turnover and replacement alone.

turnover and decreasing the quality of bureaucracy. We formalise these expectations with three testable hypotheses. Populist governments lead to *(i)* lower economic performance, *(ii)* higher bureaucratic turnover, and *(iii)* lower quality of bureaucracy.

The willingness to implement their commitment-type policies head first, rather than the recurrence of some specific “bad policies,” is the real cause of populists’ bad economic performance.

## Municipal Government and Populism in Italy

Italian municipalities represent a good case to study the consequences of populism for government performance and bureaucratic quality. First, the presence of several populist parties in Italy makes it less arbitrary for researchers to measure the populist attribute of candidates. Second, executive power is highly concentrated and populist mayors and executive committees have significant leeway to implement their commitment-type policies. Third, municipal bureaucrats play a central role in the administration of public policies, and mayors have large discretion over the appointment and removal of top municipal bureaucrats.

Consistently with various definitions of populism ([Mudde and Rovira Kaltwasser 2018](#)), empirical work generally identifies populist parties based on the presence of people-vs-elite rhetorical elements in electoral manifestos and other party sources (e.g., [Pauwels 2011](#); [Di Cocco and Monechi 2021](#)). However, measurement becomes more challenging when moving the focus from parties to individuals. Two strategies are available, based on individual behaviour or party identification. The suitability of these strategies rests on data availability for the former (e.g., individual speeches or campaign messages), and the presence of clearly populist parties for the latter. [Gennaro, Lecce, and Morelli \(2021\)](#) and [Hawkins et al. \(2019\)](#), for instance, measure populism for US congressional candidates and chief executives based on the analysis of speeches. However, when populism is a characteristic of political parties and there is no textual data to produce individual-level estimates, researchers can rely on candidates’ party identification and code populist politicians based on whether they identify with a populist party or not. While using party-level characteristics to infer individual-level characteristics clearly implies a loss of construct validity, we believe it is a suitable strategy for the Italian context, where several

parties are clearly considered populist and deviation from party lines is punished by populist parties (Fasone 2020).<sup>6</sup>

We obtain a list of populist parties in Italy from the PopuList Database, a collective enterprise that brings together researchers in the attempt to provide a comprehensive list of populist parties in Europe (Rooduijn et al. 2019). There are six populist parties in Italy, which ranks fourth in Europe for number of populist parties: Lega (Nord), Movimento 5 Stelle (Five-Star Movement, 5SM), Fratelli d'Italia, Forza Italia - Il Popolo della Libertà (only until 2018), and two minor parties, Liga Veneta and Lega d'Azione Meridionale. In Section A of the SI we provide qualitative information about these parties.

Italian municipal government features a directly-elected mayor, a directly-elected local council, and an executive committee appointed by the mayor, with elections held every five years. Mayoral candidates are linked to one or more party-lists of candidates to the local council. The lists mirror the political parties at the national level, although it is common for candidates to be linked to civic lists (i.e., *liste civiche*) that do not match any of the parties at the national level. Voters can cast a ballot for the mayoral candidate directly, for one of the lists linked to the candidate, or both. Votes cast for a list automatically count towards the number of votes of the mayoral candidate linked to the list. The candidate who wins a plurality is elected mayor, except for municipalities with more than 15,000 inhabitants, where the two candidates with the largest vote share run in a second round if none of the candidates obtain 50% of the votes in the first round.

Municipalities are responsible for a wide set of services, from primary schooling to local police, waste management, public roads and infrastructure, social services, and security. As a result, municipal governments enjoy large degrees of financial autonomy and they are responsible for the procurement of goods and services. Municipalities also have large bureaucratic apparatuses, accounting in 2017 for 12% of the 3.5 million employees working in public organisations in Italy.<sup>7</sup> The bureaucratic organisation of Italian municipal governments consists of two types of employees, those with managerial rank (*qualifica dirigenziale*) and those without

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<sup>6</sup>By February 19, 2021, the number of 5-Star-Movement members expelled from the party because of non-compliance with party directives amounts to 40. See <https://www.agi.it/politica/news/2021-02-19/m5s-espulsioni-parlamentari-governo-draghi-11478409/>.

<sup>7</sup>See the 2017 Census of Public Organisations, National Institute of Statistics <http://istat.it/it/censimenti-permanenti/istituzioni-pubbliche>.

managerial rank (*qualifica non dirigenziale*). For simplicity, we shall consider bureaucrats without managerial rank as rank-and-file employees and bureaucrats with managerial rank public managers. While rank-and-file employees have limited discretion on policy, public managers are responsible for the implementation of directives adopted by the executive committee, the financial and personnel management of the municipality, public service delivery and monitoring.<sup>8</sup>

Rank-and-file employees are generally hired through public competitions and with permanent contracts agreed at the national level through collective bargaining, whereas the mayor and the executive committee have larger discretion over the hiring and departures of public managers. There are three ways to hire public managers: public competitions published by the municipality, mobility across organizations, and direct appointment with temporary contracts. Municipal governments have increasingly opted for fixed-term contractual frameworks that allow them to exert a tighter control over bureaucrats, with many scholars arguing for the emergence of a “spoils system” in Italian local government (Borgonovi and Ongaro 2011). In the period we study, the average share of temporary contracts among all bureaucrats with managerial rank is 20%, increasing from 16% in 2003 to 25% in 2019.

While mayors have full discretion over temporary contracts, permanent contracts can be terminated only for serious failures. However, mayors have discretion over the allocation of managerial tasks and demotion within the government, and are therefore able to incentivise public managers to leave even when the contract is permanent.<sup>9</sup> Mayors can therefore affect turnover in municipal government through the termination of temporary contracts and (indirectly) through mobility across organizations for permanent contracts.<sup>10</sup>

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<sup>8</sup>The responsibilities of public managers are governed by the Consolidated Law on Local Government, Legislative Decree 267/2000, Article 107, and by the Consolidated Law of Public Employment, Legislative Decree 165/2001, Article 4.

<sup>9</sup>See Article 50(10), Lgs. D. 267/2000. Demotion practices have been documented in the literature by Doherty, Lewis, and Limbocker (2019), who find that presidents marginalise ideologically distant career executives by transferring responsibilities to other individuals.

<sup>10</sup>However, municipalities are not required by law to have public managers. Our data shows that 15% of municipalities have public managers, and the median number of managers in those municipalities is 4.

## Data

We assemble a rich dataset on municipal governments in Italy covering more than 20 years, from 1998 to 2020. We collect three sets of data: *i*) municipal elections, mayoral candidates, and partisan affiliation of politicians, *ii*) economic performance and public procurement data, and *iii*) bureaucratic composition.

We combine several sources of data. We obtained data on all municipal elections, mayoral candidates, and their party affiliations from the Historical Electoral Archive of the Ministry of the Interior (1989-2020). We used the Database on Local Administrators for other information on mayors (e.g., job, gender, and level of education), local councillors, and members of the executive committee (1998-2020). For our measures of performance we collected novel data on public procurement (2012-2020) from the archives of the National Anticorruption Authority, which we combined with data on economic performance from the National Institute of Statistics (2008-2019). Finally, data on the number of bureaucrats, their rank, education, type of contract, and data on hirings and departures is obtained from the Annual Account of the Italian General Accounting Office (2001-2019). This is at the same time an extremely rich and complex source of data which allows us to capture variation in bureaucratic composition of municipal governments over 20 years and across 8,069 unique municipal governments. To the best of our knowledge, it is the first time it is used in scholarly work.<sup>11</sup>

### Populist Candidates

We code populist mayoral candidates based on the political party-lists linked to the candidate. We delegate the identification of populist parties to the PopuList database ([Rooduijn et al. 2019](#)), commonly used in scholarly work and regularly peer-reviewed by more than eighty academics from different countries. PopuList codes parties as populist if they interpret the people and the elites as two antagonist entities; embrace the idea of popular sovereignty; and uphold an ideal vision of the people and a denigratory vision of elites.

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<sup>11</sup>Detailed description of the original sources is reported in Section B of the SI.

## Economic Performance

We present three measures of economic performance. First, we calculate the percentage of procurement contracts with cost overruns. Public procurement data is often used in the literature to study performance (Spenkuch, Teso, and Xu 2021; Decarolis et al. 2021). We follow this literature and compute the yearly percentage of all public contracts awarded by the municipality where payments are greater than the adjudicated costs. Since 2012, every municipality has to publish on its website a large set of information about each contract such as object, cost, duration, beneficiary, and awarding procedure (Article 1(32), Law 190/2012). This information is also sent to the Anticorruption Authority for verification and quality control, before being published in the authority’s open-access archives.

We web-scraped the archives and assembled a dataset with 1.25 million contracts worth at least 1,000 euros awarded by 5,527 municipalities from 2012 to 2020.<sup>12</sup> For each contract we create a variable equal to 1 if payments exceed the initial adjudicated cost, and 0 otherwise. We then compute the average of this metric for every municipality in any given year, obtaining a dataset of 25,628 municipality-year observations.<sup>13</sup>

The other two measures of performance capture the fiscal quality and sustainability of the administration, proxied by the accumulation and repayment of residual liabilities. Let us consider this example. In year  $t$ , municipality  $i$ ’s balance sheets report residual liabilities for 2,000 euros, consisting of the difference between committed and actual payments in the previous year  $t - 1$ . In year  $t$ , municipality  $i$  commits to payments for 13,000 euros and eventually pays 10,000 euros, thus accumulating other 3,000 euros of liabilities. Debt accumulation is computed as the ratio between current and initial liabilities ( $3,000/2,000 = 1.5$ ). Debt repayment is the ratio between disposed and accumulated liabilities, and measures the ability of municipality  $i$  to repay more debts than what it accumulates. A good fiscal performance is associated with low levels of debt accumulation and high levels of debt repayment.

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<sup>12</sup>Several contracts have typos in the reported information: 6,629 contracts have a termination date that is earlier than the starting date and 348 contracts have payments that are more than 100 times as large as the adjudicated costs. To clean the dataset, we removed these contracts.

<sup>13</sup>In Section E.5 in the SI we show our results are robust when computing the percentage of contracts where the difference between payments and costs exceeds 5% and 10% of the awarded costs.

## Bureaucratic Turnover and Quality

To measure the consequences of populism for bureaucratic quality and turnover, we focus on public managers. We do this for two reasons. First, public managers have large levels of discretion in the administration of policies and populist politicians are more likely to replace bureaucrats in strategic decision-making positions. Second, as discussed above, it is easier for politicians to fire and hire managers compared to rank-and-file employees.

We measure bureaucratic turnover in municipality  $i$  and year  $t$  as the sum of managers who leave (departures) and join (hirings) the government divided by the total number of managers in the same year. The precise metric is given by the following formula:

$$\text{Turnover}_{it} = \frac{\text{N. Departures}_{it} + \text{N. Hirings}_{it}}{\text{N. Public Managers}_{it}} \quad (1)$$

Perfect measures of quality of bureaucrats are hard to produce, for quality is a latent, multifaceted concept. We follow a long tradition of work in political economy that uses education as a proxy for the quality of politicians and interpret bureaucrats' quality as the percentage of managers with a university degree (in the Italian context, see e.g. [Galasso and Nannicini 2011](#); [Baltrunaite et al. 2014](#)). By jointly looking at turnover and education we provide a comprehensive picture of the costs of populism for the bureaucracy. In fact, while turnover alone can produce leadership vacuums and undermine institutional continuity in planning and implementation ([Lewis 2011](#)), these negative consequences are likely to be enhanced when coupled with a non-expert-replace-expert dynamic.

Table 1 reports descriptive statistics for the main variables in the full dataset and in the analysis dataset, namely the subset of municipalities where a populist candidate ran for office. 3,191 out of the 8,254 municipalities in our dataset had a populist candidate in any of the elections they held. In the third and fourth columns, we report statistics for municipalities where the populist candidate won and lost the elections.

Variable	Full Dataset	Analysis Dataset	Elected Mayor	
			Non-Populist	Populist
N. Observations	182,435	31,541	23,024	8,517
N. Unique Municipalities	8,254	3,191	2,951	1,229
<b>Economic Performance</b>				
Debt Accumulation	0.60	0.59	0.60	0.56
Debt Repayment	1.11	1.03	1.03	1.04
% Cost Overrun	0.07	0.09	0.09	0.10
<b>Quality of Bureaucrats</b>				
Managers' Turnover	0.24	0.23	0.23	0.23
% Graduate Managers	0.48	0.44	0.45	0.41
<b>Mayors</b>				
Age	51.82	50.55	50.31	51.10
% Graduate	0.44	0.52	0.51	0.55
% Secondary Education	0.87	0.91	0.91	0.92
% Female	0.11	0.12	0.13	0.11
% Incumbents	0.35	0.33	0.36	0.25
% White Collar	0.28	0.32	0.29	0.41
<b>Municipalities</b>				
Resident Population	7,156	20,818	18,647	26,557
Surface (sq.km)	37.19	49.58	46.44	58.01
% in Northern Regions	0.56	0.68	0.69	0.65

Table 1: Descriptive statistics of main variables in the entire dataset and in three sub-samples: municipalities where a populist candidate ran for office (Analysis Dataset), and where a populist candidate won and lost the elections.

## Close-election Regression Discontinuity Design

Identifying the effect of populist governments on our outcomes of interest is a challenging task. Municipalities governed by a populist may differ from municipalities governed by a non-populist mayor due to many unobservable characteristics. However, municipalities where populist candidates win the elections by very thin margins can be, in expectation, comparable to municipalities where the populist candidate barely lost. Close-election sharp RDD is a common method for estimating the local average treatment effect of winning candidates' characteristics on downstream outcomes. This approach has been used in the context of Italian municipalities by [Gagliarducci and Paserman \(2012\)](#) and [Casarico, Lattanzio, and Profeta \(2021\)](#), to estimate the effect of electing a female mayor on government termination and fiscal policy, by [Romarri \(2020\)](#), to estimate the effect of electing a far-right mayor on hate crimes, and by [Bordignon and Colussi \(2020\)](#), to estimate the effect of a populist candidate reaching the second round of the elections on voters' turnout and descriptive representation.

Formally, let  $E$  be a set of municipal elections in which *one* populist candidate runs against

one or more non-populist candidates. For each  $E_i$ , let  $M_i$  be the margin of victory of the populist candidate, calculated as the difference between the vote share of the populist candidate and the most voted non-populist candidate. Let  $V_i$  be a dichotomous variable equal to 1 if  $M_i > 0$  (the populist candidate wins) and 0 otherwise.  $Y_i$  is the outcome of interest. We can then define the estimand as  $\tau = \lim_{\epsilon \downarrow 0} \mathbb{E}[Y_i | M_i = \epsilon] - \lim_{\epsilon \uparrow 0} \mathbb{E}[Y_i | M_i = -\epsilon]$  as the local average treatment effect (LATE) of electing a populist candidate.<sup>14</sup>

We intentionally define  $\tau$  as the LATE of electing a populist candidate, instead of the LATE of populism alone. In fact, because close elections do not randomly assign candidate characteristics (Sekhon and Titiunik 2012), the RD estimator will recover the effect of the populist attribute of mayors plus all other individual- and municipality-level characteristics that distinguish populist from non-populist and that allow the former to remain in close-election (Marshall 2022).<sup>15</sup> In the following section we describe how balance tests for mayor-level characteristics can help us characterise the compound nature of the estimand, as recommended by recent methodological literature (Marshall 2022).

We estimate  $\tau$  with a continuity-based approach that uses nonparametric local polynomial methods for estimation and inference. We fit local WLS models where weights are determined by the triangular kernel function based on the ratio between the distance of unit  $i$  from the cutoff  $m$  and the mean-squared-error minimising bandwidth  $h$ . The closer the units are to the cutoff, the larger the weight. Units outside the optimal bandwidth receive a weight equal to zero, therefore estimation is performed on a restricted sample of units so that  $M_i \in [-h, +h]$ . We use the automatic bandwidth selector proposed by Calonico, Cattaneo, and Titiunik (2014), which aims to minimise the mean-squared-error of the local polynomial RD point estimator. Inference adjusts for the variability introduced in the bias-estimation step and uses a variance estimator that yield robust bias-corrected confidence intervals and p-values (Cattaneo, Idrobo, and Titiunik 2019). Therefore, point estimates will not be centred in the confidence interval. We estimate the following full treatment-interaction model:

$$Y_{it} = \beta V_{iT} + \phi M_{iT} + \eta V_{iT} \times M_{iT} + \zeta X_{it} + \delta_t + \gamma_T + u_{it} \quad (2)$$

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<sup>14</sup>We summarise the composition of the treated and control groups in Table C.2 in the SI.

<sup>15</sup>Similarly, Hall (2015) notes that he estimates the effect of *extremist* candidates and not that of *extremism*, and Bucchianeri (2018) identifies the effect of nominating a female *candidate* and not that of *gender* alone.

We use subscript  $T$  for election year and  $t$  for calendar year. Recall that  $V_{iT}$  is a dummy for treated units above the cutoff,  $M_{iT}$  is the margin of victory (i.e., the running variable), and  $u_{it}$  a robust error term clustered by municipality. We include a set of pre-treatment covariates  $X_{it}$  in the specification as well as year and election-year dummies ( $\delta_t$  and  $\gamma_T$ ) to boost efficiency (Calonico et al. 2019).<sup>16</sup> The coefficient  $\beta$  is the RDD estimator and identifies the average outcome jump at the cutoff after partialling out the effect of the covariates. Mayors stay in office for five years, while the outcome variables are at the municipality-calendar year level.  $\beta$  is therefore the average yearly effect of electing a populist mayor within the government term.

## Validity

An important falsification test for the RDD consists of ensuring that, near the cutoff, treated units are similar to control units in terms of observable characteristics. Balance tests are generally employed to provide evidence in support of the continuity of potential outcomes assumption, where several pre-treatment covariates are used as placebo outcomes. However, as we are interested in the effect of a candidate characteristic, the RDD estimand should be interpreted as a compound treatment effect. In most cases, this interpretation is sensible, for causal claims about fixed characteristics like populism, gender, and race should be operationalised as a “bundle of sticks” (Hall 2015; Sen and Wasow 2016). However, detecting meaningful discontinuities in candidate-level characteristics that are conceptually different from populism can help characterise the compound nature of treatment (Marshall 2022).

Figure D.2 in the SI shows balance tests for 18 pre-treatment covariates. Municipalities above and below the cutoff are very similar with respect to demographic, geographic, and political characteristics of the municipality. Importantly, we find no discontinuity for the value of the margin of victory from the previous election. Populists and non-populists mayors are also similar with respect to several individual characteristics (e.g., education, gender, and age). However, populist candidates are more likely to be new entrants compared to their non-populist counterparts. We find that the probability of populist mayors to be incumbents is 11 percentage points lower. While the absence of discontinuities in most of the covariates strengthen our confidence that municipalities where a populist mayor won and lost by thin margins are similar

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<sup>16</sup>The results do not change when adding unit dummies (see Table E.6 in the SI).

in expectation, the discontinuities in the incumbency status of non-populist mayors might suggest that populist candidates benefit electorally from being challengers, and their lack of government experience allows them to remain in close elections. This difference might confound the effect of the populist attribute of mayors. In fact, bureaucratic turnover might be the result of change in political leadership of the municipality and poorer performance might be the product of inexperience. In Section *Challenger Status and (Right-Wing) Populism* we address this source of confounding.

We perform several additional tests and we report the results in the SI. We document the absence of sorting at the cutoff with density tests aimed at detecting whether there is a proportional number of elections where populist candidates barely won or lost (Figure D.1). In Figure D.3 we report estimates from alternative placebo margins of victory ranging from  $-25\%$  to  $+25\%$ . When adjusting the estimation for multiple testing across each outcome variable, we find no discontinuities at 95% level in 77% of the tests with placebo cutoffs. While the number of discontinuities at placebo cutoffs is not negligible, we never find a simultaneous discontinuity in all our five (nor even four) outcomes in any of the 46 placebo cutoffs used. Because our measure of populism includes several different parties, discontinuities at placebo cutoffs greater than the true 0-cutoff might arise from comparing a mayor supported by a populist party with mayors supported by another populist party. Next, we perform a set of falsification tests with lagged values of the outcomes, using the margin of victory in election  $T$  to estimate jumps at the cutoff in the outcomes during the years between election  $T$  and  $T - 1$  and find a discontinuity only for one outcome (see Table D.3). We also show that the estimates are robust to alternative bandwidth selections and bandwidth selectors (Figure E.4 and Figure E.5). Finally, in Section E.2 we perform power analysis to ensure we have enough statistical power to detect the treatment effects we actually estimate (Stommes, Aronow, and Sävje 2021).

## Results

Merging the database on government performance and bureaucratic composition with election data where one populist candidate was running against at least one non-populist candidate,

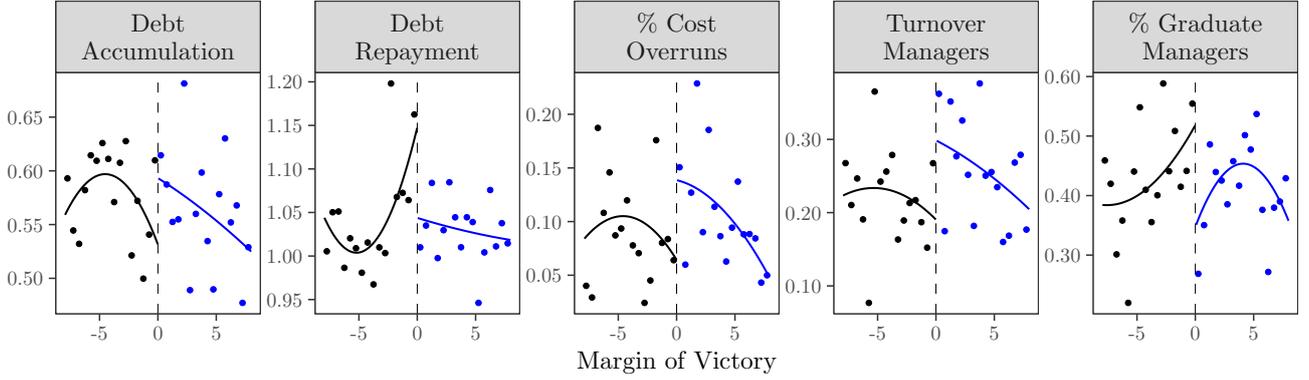


Figure 1: Binned averages of the outcomes in close electoral races. The solid lines are second-order polynomial fit using control and treated units separately. Scatter points are averaged over 0.5% margin of victory.

we obtain a sample of 7,897 municipality-elections pairs, for a total of 31,541 municipality-year observations. 2,164 elections were won by a populist (27% of the total). In most cases, populist candidates were supported by one populist party (i.e., 6,153 elections), 1,575 candidates by two, and only 169 candidates were supported by three populist parties (i.e., the right-wing populist coalition: Forza Italia, Lega (Nord), and Fratelli d'Italia).<sup>17</sup>

To visually display the discontinuities at the cutoff, Figure 1 shows binned averages of the outcome variables as a function of the margin of victory of the populist candidate. The jumps when the margin of victory equals 0 are suggestive of the expected effect of electing a populist mayor: units above the cutoff accumulate more debts, they repay a smaller fraction of accumulated debts, and they incur in more contracts with cost overruns. There are also sizeable jumps for managers' turnover and level of education that suggest turnover increases and the quality of bureaucrats decreases with the election of a populist mayor.

Table 2 shows the RD results. We report the estimated effect of electing a populist mayor *in close electoral races* on the five outcomes as well as 95% cluster-robust confidence intervals.<sup>18</sup>

Consistently with the RD plots, there are economically and statistically significant effects on government economic performance. Populists barely winning the elections accumulate more debts, with the ratio of new over old liabilities increasing by 0.03 points, equal to 5% the average value in the dataset (i.e., 0.59). Municipalities with a populist mayor also do a worse job of repaying debts accumulated in the current year, with debt repayment decreasing by

<sup>17</sup>In Table A.1 in the SI we show the distribution of populist candidates across each pair of supporting populist parties.

<sup>18</sup>Results without covariates are reported in Table E.5 in the SI. Only for debt accumulation the 95% CI includes 0 (robust p.value = 0.348).

<i>Outcomes</i>	<b>Economic Performance</b>			<b>Quality of Bureaucrats</b>	
	Debt Accumulation	Debt Repayment	% Cost Overruns	Turnover Managers	% Graduate Managers
<i>Estimate</i>	0.032	-0.046	0.048	0.130	-0.131
<i>95% CI</i>	[0.015, 0.058]	[-0.081, -0.021]	[0.009, 0.102]	[0.085, 0.199]	[-0.194, -0.093]
<i>h</i>	9.639	13.264	12.882	8.340	6.208
<i>Obs. Used</i>	3,523	4,746	1,232	1,449	1,000

Table 2: RD estimates constructed using local polynomial estimators with triangular kernel. Robust 95% confidence interval constructed using bias-correction with cluster robust standard errors at municipality level,  $h$  is the MSE-optimal bandwidth. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year and year-election dummies. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019. Analysis implemented with the *rdrobust* package in R (Calonico, Cattaneo, and Titiunik 2015).

-0.05 points (-4% compared to the mean in the data), and the share of procurement contracts with cost overruns increases by 4.8 percentage points. Electing a populist mayor also leads to an increase in turnover among public managers by 0.13 points (average turnover is 0.23), and a simultaneous decrease in the percentage of managers with a university degree by -13.1 percentage points.<sup>19</sup>

In Section F.3 in the SI we perform additional tests to detect heterogeneity in treatment effects based on the first or second part of government term, geographical location, and population size of municipalities and we do not find marked differences, except for the effects on performance outcomes in the sample of municipalities with population size below the median value, which generally display larger effects.

### Forced or Voluntary Departures

Our expectation for which populist mayors get rid of expert managers with detrimental consequences for the quality of the bureaucracy hinges on expert bureaucrats being forced to leave the administration. However, an alternative, observationally equivalent mechanism could be at play, whereby managers voluntarily decide to leave the administration following the election of a populist mayor. For instance, Bolton, Figueiredo, and Lewis (2021) argue that top bureaucrats might decide to leave the organisation when facing a newly-elected principal

<sup>19</sup>We find positive effects on both the number of hirings and departures (see SI Table F.8). Moreover, as a falsification test, we find much smaller effects for rank-and-file employees, with the effect on turnover even being of opposite sign (see Table F.9).

Departure Category	Forced	Voluntary	Total
Other Reasons	?	?	2,341
Resignation		X	1,800
Transfer to Other Administrations		X	863
Forced Retirement for Age Limits	X		763
Contract Termination	X		182
Contract Termination, 40 years of contributions			71
Firing	X		60
Transfer due to Outsourcing			25

Table 3: Managers’ departure categories and number of instances in the data (2001-2019).

with diverging policy positions, either because they are marginalised by the new government or because the value they obtain from public office decreases.

To adjudicate between these two mechanisms, we dig into the data on bureaucrats’ departures, which allows us to code whether the departure was a mayor’s or bureaucrat’s decision. Our data features eight categories of departures, which are displayed in Table 3 and described in more detail in Section F.2 in the SI. After reviewing the legislation with the support of a public manager from an Italian municipality, we are able to classify categories based on whether they represent a decision of the mayor (forced) or of the manager (voluntary). Some categories are either too ambiguous to be considered forced or voluntary departures (e.g., “Other Reasons”) or they are forms of natural termination of the contract (e.g., “Termination of Contract (40 years of contributions)”).

We then estimate the effect of electing a populist mayor on four separate outcomes: total departures, voluntary and forced departures, and “other reasons,” all divided by the total number of managers. It should be acknowledged that the large number of “Other Reasons” imposes caution in the interpretation of this test, for we do not know which category is more likely to pass under the other-reasons umbrella. Figure 2 displays the results. First, consistently with the increased turnover, we show that the number of departures increases with populist mayors. Despite the largest effect being estimated for the “Other Reasons” category, the positive effect on forced departures – coupled with the approximately 0-effect on voluntary departures – lends support to the proposed mechanism.

These results corroborate the expectations we derived from characterising populism as a form of unconditional policy commitment. Populist governments incur in larger debts and inef-

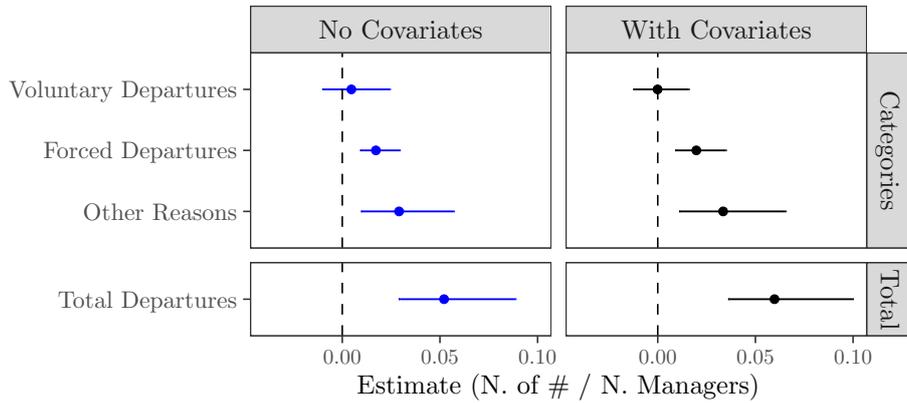


Figure 2: RD estimates with 95% cluster-robust CI of the effect of electing a populist mayor on voluntary and forced departures as percentages of the number of managers. Same estimation, covariates, and period of analysis reported in Table 2.

iciencies in managing procurement contracts. In the attempt to get rid of internal bureaucratic resistance, populists are also more likely than non-populist to replace expert bureaucrats with less educated ones, thus decreasing the overall quality of the bureaucracy. We therefore interpret lower performance and weakened bureaucracy as two sides of the same coin: populists’ costly commitments.

## Challenger Status and (Righ-Wing) Populism

We present two robustness tests to further strengthen our results. First, we address the possibility of the estimated effects being confounded by the challenger status of populist mayors. The lower probability of populist mayors to be incumbent might confound the effect of the populist attribute of mayors. Changes in performance and bureaucratic organisation might be the result of inexperience and reasonable change in administrative leadership resulting from the election of a new mayor (see e.g., Bolton, Figueiredo, and Lewis 2021). To rule out this possibility, we perform two additional tests removing *i*) incumbent mayors (i.e., mayors serving the second term), as well as *ii*) mayors who are serving the first term (i.e., challengers) but who are supported by an incumbent party. Specifically, we remove incumbent mayors and first-term mayor if at least one of the parties supporting the mayor was also supporting the previous, incumbent mayor. With these two tests we therefore address both candidate- and party-level incumbency status.

Figure 3 shows the estimated effects across the total sample of mayors, the sample where we

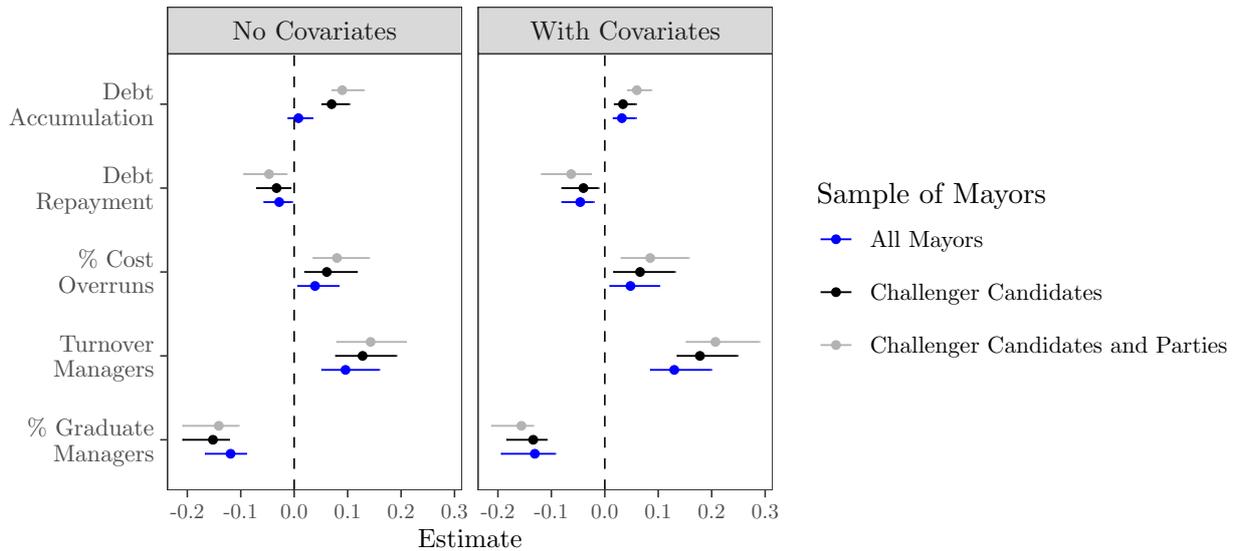


Figure 3: RD estimates of the effects of electing a populist mayor excluding incumbent mayors (we keep the first term of all mayors who have been elected more than once) and excluding incumbent mayors and non-incumbent mayors supported by at least one incumbent party. Same estimation, covariates, and period of analysis reported in Table 2.

compare populist and non-populist challengers, and a sample where we compare populist and non-populist challengers who are not supported by incumbent parties. The estimates remain in the expected direction and distinguishable from 0 across all the two additional outcomes, suggesting that the populist attribute of the mayors – rather than their challenger status – is driving the results.

Second, except for the 5SM, the other populist parties are also right-wing, hence there may be a concern about what is driving the results, whether being populist or being right-wing (Hunger and Paxton 2021). While it is not possible to conclusively isolate the effects of the two components, we show that the results are not exclusively driven by one single right-wing populist party. In Figure 4 we display the main results reported in Table 2 sequentially moving right-wing populist parties (reported in the panel labels) to the control group, therefore inflating the control group with right-wing candidates. Because of some disagreement in the literature on whether Berlusconi’s parties should be coded as populist (Van Kessel 2015, 54), we also report the results after coding *both* Forza Italia and Il Popolo della Libertà as non-populist. All the estimated effects are in the same direction as the main results and statistically significant at 95% (only one at 90% level).

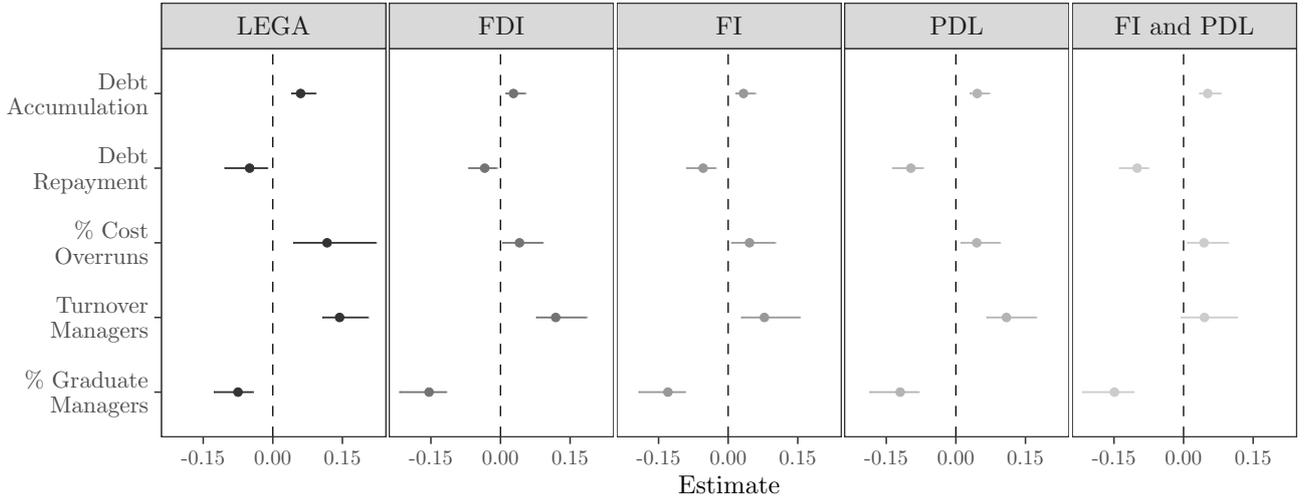


Figure 4: RD estimates with 95% cluster-robust CI of the effect of electing a populist mayor under alternative coding strategies. Panel labels report the party omitted when coding the candidates as populist. Same estimation, covariates, and period of analysis reported in Table 2.

## Conclusions

In this paper we explore how populism affects government economic performance and bureaucratic quality. We take stock of different theoretical approaches to the study of policy-making under populism, which argue that when citizens lose trust in traditional political parties and elites, populist politicians have an incentive to propose a commitment-type policy platform that can be easily monitored by voters. We derive testable implications of populists' commitment-type policies, and we test them in the context of Italian municipalities, where mayors have large influence on policy-making and bureaucratic organisation. We analyse all municipal elections in which a populist mayor runs for office in a regression discontinuity design and find that when a populist mayor barely wins the election, debts and cost overruns on procurement contracts increase, suggesting lower financial sustainability and efficiency in government economic performance. We also find that populist mayors lead to higher turnover among top bureaucrats, who in turn are on average less educated. We present several robustness tests that reinforce our confidence in the commensurability of our estimates with the effect of electing a populist mayor in close electoral races, disentangling the effect of the populist attribute of mayors from their challenger status.

Three notes on the generalisability of our results are in order, concerning the representativeness of our institutional context, the scope conditions of the theoretical set-up, and the

outcomes we selected for this study. First, our study of municipal government raises questions about the implications of our findings for higher levels of government. While we study populism in relatively small political communities, there are two reasons why we expect our findings to travel to higher levels of government. First, it has been shown in the literature that populist governments can influence policy at the local (Paxton 2019) as well as national level (Funke, Schularick, and Trebesch 2020; Biard, Bernhard, and Beta 2019). Second, the cost of populists' commitments might be larger where politicians have more influence over more salient policies. The policies implemented by populists controlling regional or sub-national governments, large ministries, or public organisations in charge of the administration of large programmes might have even larger effects on performance and bureaucratic organisation.

At the same time, the theory discussed in this paper is suitable for explaining the policy consequences of populism in political systems with low levels of power sharing and considerable influence of elected officials on bureaucratic bodies and policy outcomes. Coalition governments including populist parties might significantly constrain the populist agenda, and inter-party bargaining can marginalise the demands of populist parties. Similarly, political systems which do not grant considerable discretion to chief executives might pose institutional constraints on the commitment-type agenda of populists which would mitigate the consequences of populist government. It seems reasonable for this logic to apply for local government too. For instance, Paxton (2019) finds that the populist mayor of the city of Padua in Italy managed to implement restrictive anti-immigrants policy as a result of the large formal power that mayors enjoy, whereas a feeble influence was achieved by the populist mayor of the city of Wels, in Austria, where municipal governments enjoy a lower degree of autonomy. On the bureaucracy side, our framework applies to bureaucracies with clear tasks and organisations, which are delegated significant stocks of discretion in managing policies. In political systems where crony or corrupted practices prevail it is harder for expertise to affirm as a distinctive feature of bureaucratic policy-making, and populist politicians might not be concerned about replacing experts with non-experts.

Finally, it should be noted that the effects we find on economic performance might be only one fraction of the factors affected by populist government. Longer-term effects might hamper the economic attractiveness of the municipality with detrimental consequences on investments

and local economic growth (in a similar vein to the findings in [Funke, Schularick, and Trebesch 2020](#)). On the bureaucracy side, while the data allows us to prioritise the demand-side mechanism of turnover over a supply-side one based on bureaucrats' self-selection and voluntary departures, we cannot rule out alternative ways through which populism undermines bureaucracy. An important one, highlighted in theoretical work, is that those expert bureaucrats who remain in the administration can “pause” their commitment to good-quality policies and feign to be non-expert while waiting out the incumbent government ([Sasso and Morelli 2021](#); [Cameron and Figueiredo 2020](#)). Future research could study other facets of performance and examine the conditions under which bureaucrats are willing to compromise on policy today to remain in their post tomorrow.

While it is important to highlight these specificities of the Italian context and more general scope conditions of the theoretical framework, we believe that the main gist of the paper is rather general. Populism is on the rise across the world<sup>20</sup> and it is likely to have sizeable consequences for the performance of the government and interfere with the appointment and removal decisions that characterise the relationship between political principals and bureaucratic agents.

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<sup>20</sup>The aggregate 2020 vote share of populist parties in Europe doubled with respect to the total share in 2010 (from approximately 15% to more than 30%) ([Rooduijn et al. 2019](#)).

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# Supplemental Information

## Contents

A	Populist Parties in Italy . . . . .	2
A.1	Distribution of Candidates Across Supporting Populist Parties . . . . .	2
B	Datasets . . . . .	3
C	Treatment and Control Group . . . . .	4
D	Validity of RDD . . . . .	5
D.1	Continuity of Density . . . . .	5
D.2	Continuity of Potential Outcomes . . . . .	6
E	Robustness Tests . . . . .	9
E.1	Alternative Bandwidths and Bandwidth Selectors . . . . .	9
E.2	Power Analysis . . . . .	11
E.3	Main Results without Covariate-Adjustment . . . . .	12
E.4	Main Results with Unit Dummies . . . . .	13
E.5	Alternative Measures of Cost Overruns . . . . .	14
F	Additional Evidence on the Mechanism . . . . .	15
F.1	Unpacking the Effect on Turnover . . . . .	15
F.2	Supply and Demand side of Turnover . . . . .	16
F.3	Heterogeneity Analysis . . . . .	17

## A Populist Parties in Italy

In this section we provide qualitative information about the five main populist parties in Italy.

The Lega (Nord) was established in 1991, with a strong regionalist and secessionist focus (Meardi and Guardiancich 2021). Together with Giorgia Meloni’s Fratelli d’Italia, established in 2012, the new leadership of Salvini in 2013 moved the party to a more right-wing, conservative, nationalist platform, with a strong focus on anti-immigration and security. The Five-Star Movement, instead, was established as a more participatory and deliberative platform, advocate of direct democracy, and with a strong anti-elite identity (Mosca and Tronconi 2019). Despite some stark differences, they all share the aversion to austerity measures and see the European Union as a technocratic organisation which puts at risk the national interest (Caiani and Graziano 2019). Silvio Berlusconi’s Forza Italia and, then, Il Popolo della Libertà, has been portrayed as the first instance of populist party in Italy, known for its anti-establishment and people-centric rhetoric against corrupt elites, intellectuals and members of the judiciary (Van Kessel 2015). However, the populist nature of Forza Italia, which replaced Il Popolo della Libertà after its dissolution in 2013, has been recently questioned in the literature. PopuList, for instance, stops coding Forza Italia as populist after 2018, for it is considered a borderline case (Rooduijn et al. 2019). In Figure 4 we show that our results are robust to omitting these two parties from the list of populist parties.

The features that can be found in Italian populism are far from being a distinct characteristic of Italian politics. The combination of anti-establishment sentiment with right-wing conservatism, holistic-representation ambitions, and economic nationalism can be found in many countries and regions of the world, from Latin to North America, Europe, Asian, and Africa alike (Moffitt 2016).

### A.1 Distribution of Candidates Across Supporting Populist Parties

Table A.1 shows the distribution of populist candidates across supporting populist parties for all mayors supported by one (diagonal of the matrix) or two parties (every other entry).

	Forza Italia	Pop. Libertà	Lega (Nord)	Fratelli d’Italia	Movimento 5 Stelle	Liga Veneta	Lega Az. Meridionale
Forza Italia	1,310	2	352	85	0	0	0
Popolo della Libertà		781	738	13	0	0	0
Lega (Nord)			3,211	384	0	1	0
Fratelli d’Italia				84	0	0	0
Movimento 5 Stelle					764	0	0
Liga Veneta						3	0
Lega d’Azione Meridionale							0

Table A.1: Distribution of populist candidates across supporting populist parties. Each entry in the matrix represent the number of candidates supported by the parties in the respective row and column. Diagonal of the matrix shows the number of candidates supported by one populist parties. The 169 instances where mayors are supported by the right-wing populist coalition (Forza Italia, Lega (Nord), and Fratelli d’Italia) have been omitted.

## B Datasets

In this section we provide detailed information on the source of data used to assemble the dataset.

We obtained election data directly from the Ministry of Interior. Similar data – with less detailed information on the denomination of the party-lists supporting each candidate – is reported in the Historical Electoral Archive accessible at <https://elezioni.interno.gov.it/opendata>. The dataset includes information about every mayoral candidate in the total population of municipal elections from 1989 to 2020. We focus on the elections where *one* populist candidate ran against *at least one* non-populist candidate. Municipalities with more than 15,000 inhabitants have a two-round electoral system, where the two most voted candidates compete in a second round when no one obtains more than 50% of votes in the first round. When a second-round occurred, we focused on the two candidates running in the second round. As a result, if a populist was running in the first round but did not qualify to the second round, the election is excluded from the sample.

Data on bureaucratic composition of municipal governments is obtained from the Annual Account of the Italian General Accounting Office, which is a department within the Ministry of the Economy and Finance (available at <http://contoannuale.mef.gov.it>). The richness of this data allows us build fine-grained measures of bureaucratic turnover and level of education from 2001 to 2019 across all Italian municipalities. Importantly, we are able to focus on key bureaucrats within municipal governments, namely those with managerial rank. To do this, we subset each datasets of the Annual Account to macro-categories of contracts which contain the word *dirigente* (manager).

Database on Local Administrators with information on the number of local councillors and members of the executive committee, as well as the job, gender, date and place of birth of the mayor, is available at <http://dait.interno.gov.it/elezioni/open-data>. Data on debt repayment for the period 2008-2019 available at <http://dati.statistiche-pa.it>. Procurement data available at <https://dati.anticorruzione.it>. Demographic and territorial data were downloaded from the National Institute of Statistics database at <http://dati.istat.it/>.

Because not every dataset resorts to unique code identifiers, we alternated merging strategies using strings that combined both the municipality and region name, the unique identifiers assigned by the National Institute of Statistics, or the unique code attached to each municipality's budget data.

## C Treatment and Control Group

In the table below we report a summary description of the treatment and control groups for the RD analysis.

---

### Treatment Group

- 1) one populist candidate runs against one or more non-populist candidates in election year  $T$
- 2) the populist candidate barely wins
- 3) we analyse outcomes for that municipality during every calendar year  $t$  until the next election in  $T + 1$

---

### Control Group

- 1) one populist candidate runs against one or more non-populist candidates in election year  $T$
  - 2) the populist candidate barely loses
  - 3) we analyse outcomes for that municipality during every calendar year  $t$  until the next election in  $T + 1$
- 

Table C.2: Description of treatment and control groups.

## D Validity of RDD

In this section we report falsification tests for the regression discontinuity design.

### D.1 Continuity of Density

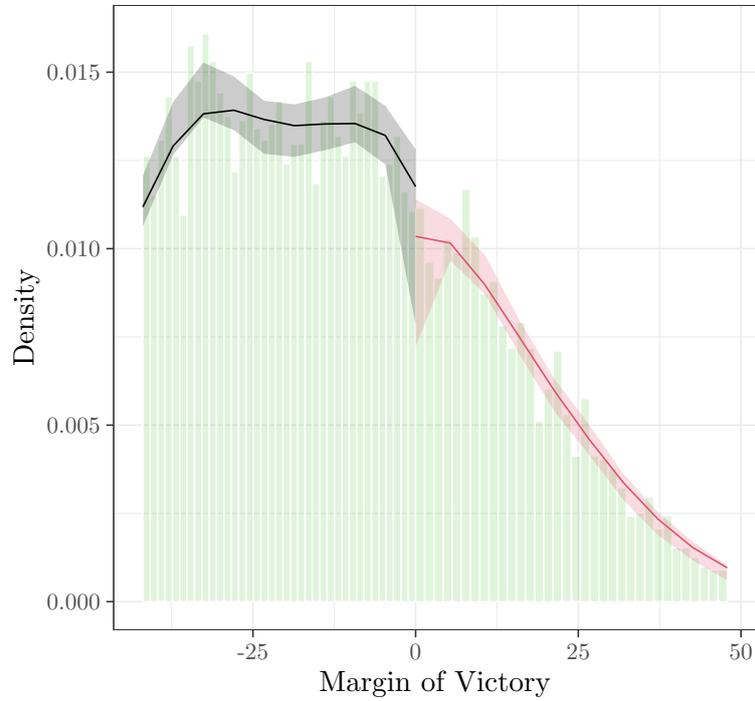


Figure D.1: Manipulation test using the local polynomial density estimator proposed in [Cattaneo, Jansson, and Ma \(2020\)](#). Histogram estimate of the running variable computed with default values in R; local polynomial density estimate (solid dark and red) and robust bias corrected confidence intervals (shaded dark and red) computed using *rddensity* package in R. We cannot reject the null hypothesis of no sorting with  $p.value = 0.547$ .

## D.2 Continuity of Potential Outcomes

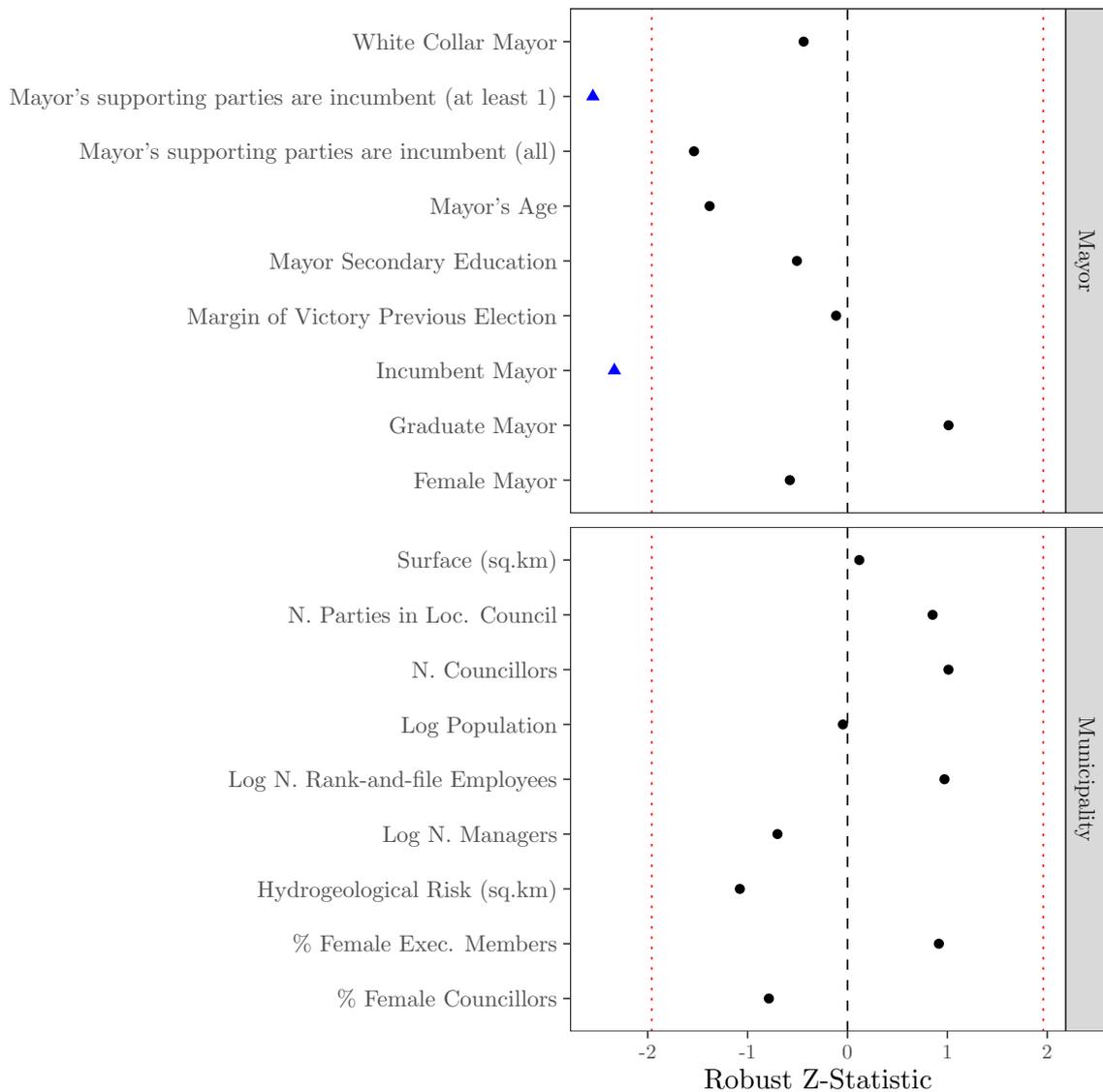


Figure D.2: RD robust Z-statistics of the effect of electing a populist mayor on pre-treatment covariates with vertical line at 95% confidence level. Unit of analysis is municipality-election year for covariates that do not change within the government term, whereas unit of analysis is municipality-calendar year for covariates that do change within the government term (i.e., population, number of rank-and-file employees, and n. of public managers). Estimates constructed using local polynomial estimators with triangular kernel and CER-optimal bandwidth (as suggested by Cattaneo, Idrobo, and Titiunik 2019, Ch. 5). Robust p-values using bias-correction with cluster robust standard errors at municipality level. Same covariates as in Tables 2, except for the covariate used as outcome variables (year dummies included only for covariates that change at calendar-year level). Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

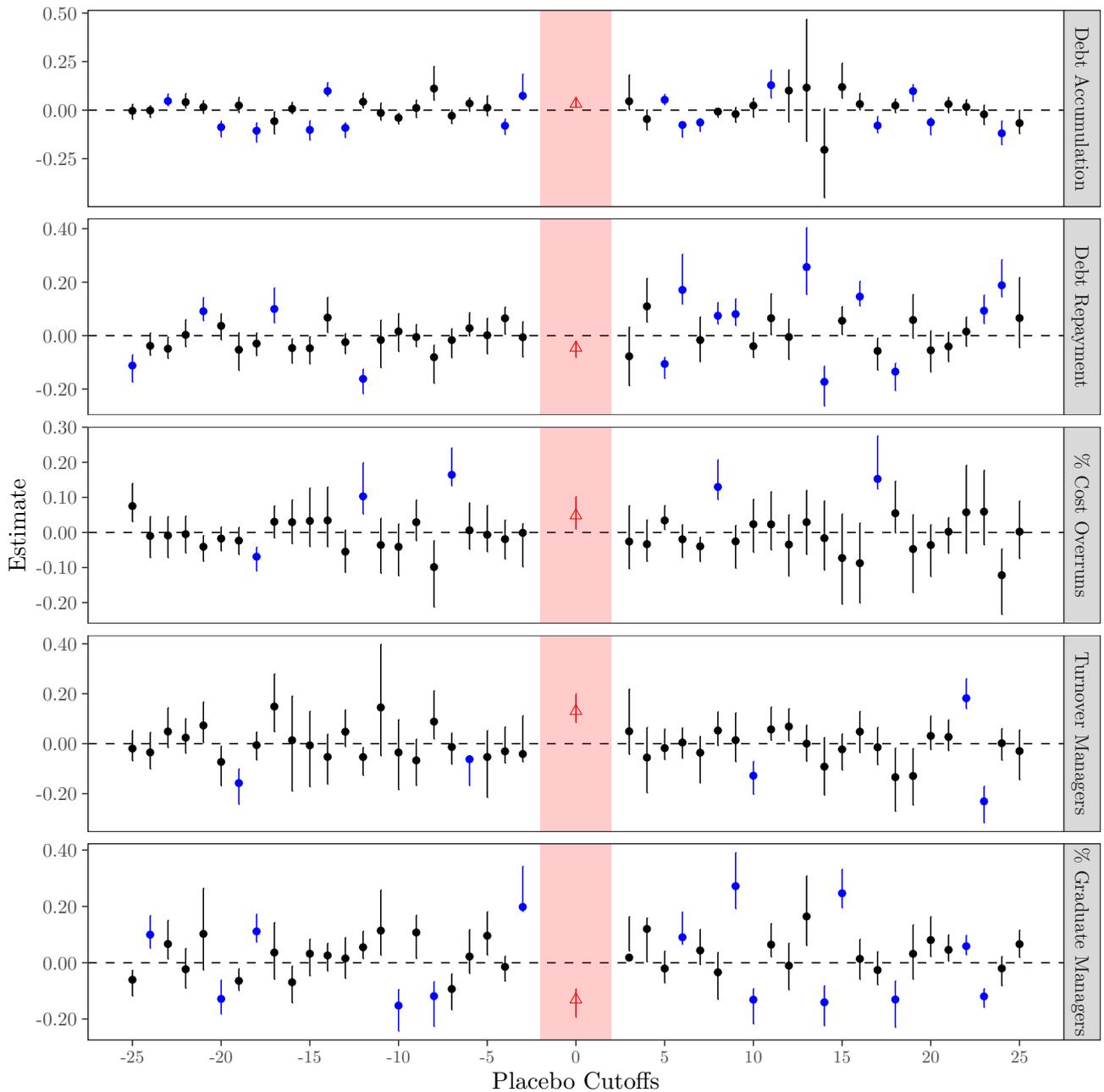


Figure D.3: RD estimates with 95% cluster-robust CI of the effect of electing a populist mayor with placebo cutoffs. Red coefficients at true cutoff (margin of victory = 0). Blue coefficients when p.value after multiple testing adjustment is smaller than 0.05. Multiple-testing adjustment performed separately for each outcome variable with Bonferroni procedure to control for the false discovery rate. RD estimates constructed separately on control unit when placebo cutoff < 0, and on treated unit when placebo cutoff > 0. Placebo cutoffs very close to 0 (i.e., 1%, 2%) omitted due to small sample size. Estimation performed using local polynomial estimators with triangular kernel and MSE-optimal bandwidth. Confidence interval constructed using bias-correction with cluster robust standard errors at municipality level. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year and year-election dummies. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019. We fail to detect a discontinuity statistically significant effects in 77% of the tests.

<b>Lagged DV</b>					
<i>Outcomes</i>	<b>Economic Performance</b>			<b>Quality of Bureaucrats</b>	
	Debt Accumulation	Debt Repayment	% Cost Overruns	Turnover Managers	% Graduate Managers
<i>Estimate</i>	-0.004	0.003	0.014	-0.034	-0.090
<i>95% CI</i>	[-0.036, 0.038]	[-0.049, 0.067]	[-0.029, 0.064]	[-0.099, 0.016]	[-0.164, -0.041]
<i>h</i>	18.420	13.348	16.720	15.314	9.742
<i>Obs. Used</i>	3,513	2,637	416	1,874	1,295

Table D.3: RD estimates of the effects of electing a populist mayor in year  $T$  on outcomes observed between election  $T$  and  $T - 1$ . Estimates constructed using local polynomial estimators with triangular kernel. Robust 95% confidence interval constructed using bias-correction with cluster robust standard errors at municipality level,  $h$  is the MSE-optimal bandwidth. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year and year-election dummies. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

## E Robustness Tests

### E.1 Alternative Bandwidths and Bandwidth Selectors

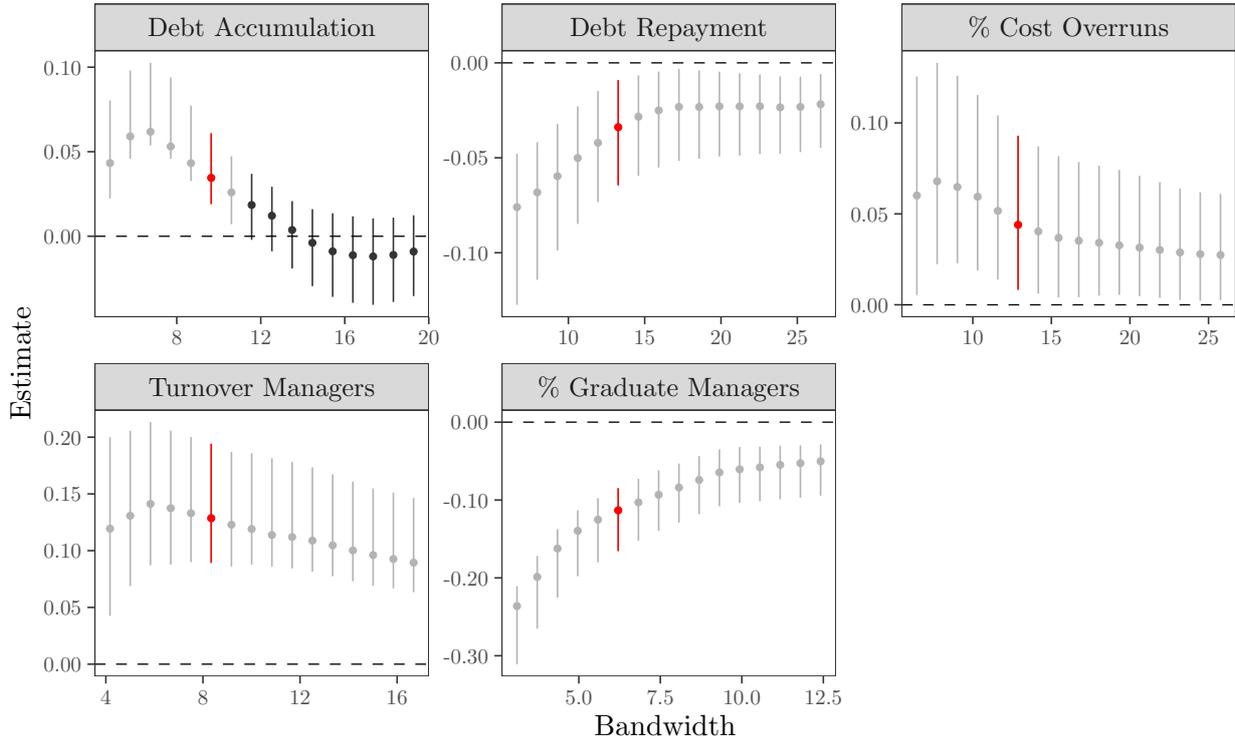


Figure E.4: RD estimates with 95% cluster-robust CI of the effect of electing a populist mayor with alternative bandwidths. Red coefficients estimated with MSE-optimal bandwidth and grey coefficients statistically significant at 95% level. Alternative bandwidths selected automatically ranging from half to 2 times the MSE-optimal bandwidth at interval of .1 (i.e.,  $h \times j$ , where  $j = 0.5, 0.6, 0.7, \dots, 2$ ). Estimates constructed using local polynomial estimators with triangular kernel and MSE-optimal bandwidth. Robust p-values and confidence interval constructed using bias-correction with cluster robust standard errors at municipality level. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year and year-election dummies. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

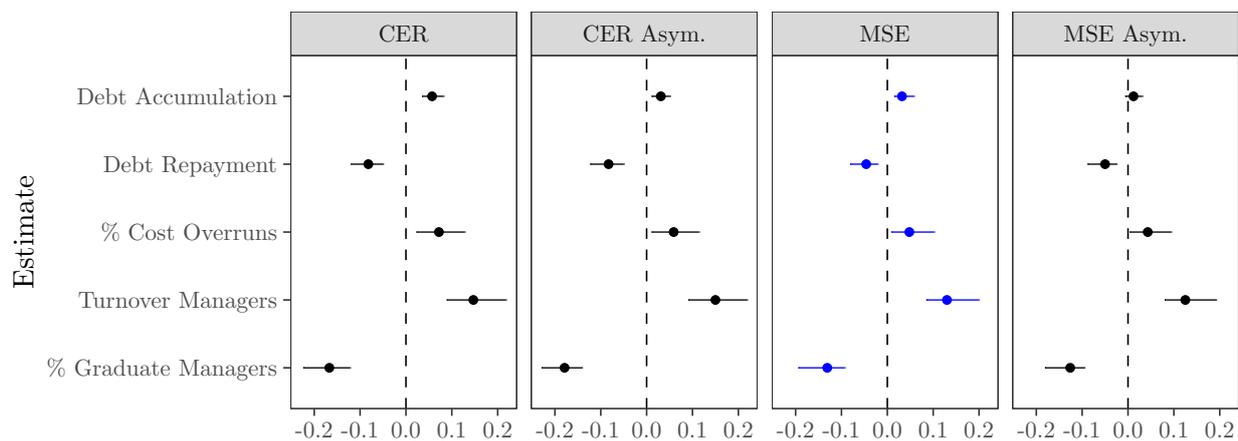


Figure E.5: RD estimates with 95% cluster-robust CI of the effect of electing a populist mayor with alternative bandwidth selection procedures (panel labels): mean-squared-error and coverage-error-rate optimal selectors both symmetric and asymmetric. MSE-optimal selector minimizes the asymptotic MSE of the point estimator, whereas the CER-optimal bandwidth minimizes the asymptotic coverage error rate of the robust bias-corrected confidence interval (Cattaneo, Idrobo, and Titiunik 2019). Blue coefficient for MSE-optimal selector used in baseline analysis. Estimates constructed using local polynomial estimators with triangular kernel. Robust p-values and confidence interval constructed using bias-correction with cluster robust standard errors at municipality level. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year and year-election dummies. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

## E.2 Power Analysis

Stommes, Aronow, and Sävje (2021) assess the validity of 36 studies using the RDD published in top political science journals and find that most studies are poorly powered to detect anything but large effects. To strengthen the validity of our results, we conducted power analysis with the method implemented in the *rdpower* package (Cattaneo, Titiunik, and Vazquez-Bare 2019). We estimate the power of a two-tailed test at the 5% significance level. We use the default settings of the package and investigate power with respect to an effect size equal to the one estimated and reported in Table 2. Table E.4 below reports the statistical power to detect such effects. Overall the probability of detecting true positive is greater than (or approximates) the conventional 0.8 threshold for each of our outcomes.

Outcome	Target Effect Size (Estimated Effect Size)	Power
Debt Accumulation	0.032	0.861
Debt Repayment	-0.046	0.967
% Cost Overruns	0.048	0.720
Turnover Managers	0.130	1.000
% Graduate Managers	-0.131	1.000

Table E.4: Statistical power achieved by an effect size equal to effects reported in Table 2. Analysis implemented with *rdpower* package in R. Calculation performed with same estimation and no covariates.

### E.3 Main Results without Covariate-Adjustment

<i>Outcomes</i>	Economic Performance			Quality of Bureaucrats	
	Debt Accumulation	Debt Repayment	% Cost Overruns	Turnover Managers	% Graduate Managers
<i>Estimate</i>	0.008	-0.028	0.039	0.096	-0.119
<i>95% CI</i>	[-0.012, 0.034]	[-0.057, -0.004]	[0.006, 0.083]	[0.051, 0.159]	[-0.167, -0.090]
<i>h</i>	13.130	15.062	14.417	9.191	7.666
<i>Obs. Used</i>	5,177	5,789	1,538	1,805	1,464

Table E.5: RD estimates constructed using local polynomial estimators with triangular kernel. Robust 95% confidence interval constructed using bias-correction with cluster robust standard errors at municipality level,  $h$  is the MSE-optimal bandwidth. No covariates included. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

## E.4 Main Results with Unit Dummies

<i>Outcomes</i>	Economic Performance			Quality of Bureaucrats	
	Debt Accumulation	Debt Repayment	% Cost Overruns	Turnover Managers	% Graduate Managers
<i>Estimate</i>	0.031	-0.044	0.049	0.124	-0.134
<i>Robust 95% CI</i>	[0.014, 0.056]	[-0.080, -0.019]	[0.010, 0.103]	[0.078, 0.193]	[-0.197, -0.095]
<i>h</i>	9.76	13.20	12.92	8.46	6.24
<i>Obs. Used</i>	3,562	4,727	1,237	1,456	1,000

Table E.6: RD estimates constructed using local polynomial estimators with triangular kernel. Robust 95% confidence intervals constructed using bias-correction with cluster robust standard errors at municipality level,  $h$  is the MSE-optimal bandwidth. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year, year-election, and *unit* dummies. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

## E.5 Alternative Measures of Cost Overruns

<i>Outcomes</i>	<b>Alternative Measures of Cost Overruns</b>		
	% Cost Overruns	% Cost Overruns (> 5%)	% Cost Overruns (> 10%)
<i>Estimate</i>	0.048	0.035	0.036
<i>Robust 95% CI</i>	[0.009, 0.102]	[0.004, 0.080]	[0.005, 0.078]
<i>h</i>	12.88	13.01	12.66
<i>Obs. Used</i>	1,232	1,243	1,214

Table E.7: RD estimates of the effect of electing a populist mayor on three alternative measures of cost overruns: percentage of procurement contracts with payments greater than awarded costs (baseline measure used in results in Table 2), and percentage of contracts with payments exceeding costs by at least 5% and 10%. Estimates constructed using local polynomial estimators with triangular kernel. Robust 95% confidence interval constructed using bias-correction with cluster robust standard errors at municipality level,  $h$  is the MSE-optimal bandwidth. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year and year-election dummies. Period of analysis 2012-2020.

## F Additional Evidence on the Mechanism

In this section we provide additional tests to clarify or corroborate the mechanism underpinning our theoretical argument.

### F.1 Unpacking the Effect on Turnover

<i>Outcomes</i>	<b>Components of Turnover</b>		
	Hirings/Total	Departures/Total	Hirings + Departures/Total
<i>Estimate</i>	0.064	0.060	0.130
<i>Robust 95% CI</i>	[0.033, 0.110]	[0.036, 0.100]	[0.085, 0.199]
<i>h</i>	9.13	9.07	8.34
<i>Obs. Used</i>	1,603	1,575	1,449

Table F.8: RD estimates of the effect of electing a populist mayor on managers' turnover and the two hirings and departures components thereof. Estimates constructed using local polynomial estimators with triangular kernel. Robust p-values and confidence interval constructed using bias-correction with cluster robust standard errors at municipality level,  $h$  is the MSE-optimal bandwidth. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year and year-election dummies. Period of analysis: 2001-2019.

<i>Outcomes</i>	<b>Rank-and-File Employees</b>	
	Turnover	% Graduate
<i>Estimate</i>	-0.021	-0.017
<i>Robust 95% CI</i>	[-0.033, -0.015]	[-0.024, -0.013]
<i>h</i>	8.41	12.73
<i>Obs. Used</i>	4,448	6,822

Table F.9: RD estimates of the effect of electing a populist mayor on turnover and education of rank-and-file employees. Estimates constructed using local polynomial estimators with triangular kernel. Robust p-values and confidence interval constructed using bias-correction with cluster robust standard errors at municipality level,  $h$  is the MSE-optimal bandwidth. Covariates include: population (log), surface (sq.km), surface at hydro-geological risk (sq.km), number of local councillors, gender, secondary education, degree, and white-collar job of mayor (all dichotomous), year, year-election. Period of analysis: 2001-2019.

## F.2 Supply and Demand side of Turnover

Here we describe the various categories of managers' departures analysed in Figure 2.

For every municipality with public managers, we have data on the number of departures. Every departure is associated with one of eight categories listed in Table 3. These categories can be clustered in three groups: voluntary departures, forced departures, and a third cluster of unknown reasons.

*Resignations* consists of public managers unilaterally deciding to terminate their contract of employment. Similarly, it is public managers who need to apply and request to be *transferred to other administrations*. On the other hand, mayors can decide to *terminate the contract of public managers* (especially with fixed term contracts), they can *fire* them, and they can *force the retirement* when there are age limits. Normal *retirements* automatically occur after 40 years of work, and *transfers due to outsourcing* materialise when public services are outsourced to a public-owned company. There is also a big black box consisting of "*other reasons*", which is the largest category of managers' departures.

### F.3 Heterogeneity Analysis

In Table F.10 below, we report the estimated differences in the RD coefficients from two samples: second and first half of the government term. There are no differences that are statistically significant at standard confidence levels. The results do not change if we perform the analysis separately for each year of the government term.

Outcome	First Half		Second Half		Difference (Second-First)		
	Estimate	SE	Estimate	SE	Difference	SE	p.value
<i>Debt Accumulation</i>	-0.030	0.044	0.014	0.027	0.044	0.052	0.394
<i>Debt Repayment</i>	0.010	0.044	-0.060	0.038	-0.070	0.058	0.229
<i>% Cost Overruns</i>	0.019	0.027	0.050	0.037	0.031	0.046	0.499
<i>Turnover Managers</i>	0.065	0.062	0.103	0.050	0.038	0.080	0.633
<i>% Graduate Managers</i>	-0.116	0.049	-0.066	0.048	0.050	0.069	0.466

Table F.10: RD estimates of the effect of electing a populist mayor estimated on two sub-samples of observations in first (first three years) and second half of the government mandate (remaining years), as well as the difference in RD estimates. Same estimation of baseline analysis. To maximise sample size, no covariates were included. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

One reason why we fail to detect significant differences over the government term is that personnel changes are spread quite evenly over the government term, although with a slight decreasing trend, as evidenced by the descriptive statistics in Table F.11 below.

Year of Government Term	Turnover		Departures		Hirings	
	Mean	SD	Mean	SD	Mean	SD
1	0.34	0.54	1.10	2.74	0.67	1.74
2	0.23	0.43	0.65	2.03	0.54	1.14
3	0.20	0.34	0.63	1.81	0.47	2.13
4	0.20	0.39	0.65	1.68	0.44	1.25
5	0.19	0.38	0.60	1.46	0.37	1.12

Table F.11: Descriptive statistics of turnover, hired and fired managers over the government term. Departures and Hirings are averages of the number of managers that leave and join the municipality. Turnover is measured as the sum of managers who join and leave divided by the total number of managers in any given year.

As for the geographical differences, in Table F.12 below we report the estimated differences in the RD coefficients from two samples: municipalities located in the Centre/South and North of the country, which is the most intuitive geographical divide given the persisting socio-economic differences between the North and South of the country.

As displayed in the table, there are two differences that are statistically significant at standard confidence levels for two performance outcomes. These differences, however, are in contrasting directions. While populists elected in the Centre-South do a better job of repaying debts, they exhibit larger shares of procurements contracts with cost overruns.

Outcome	Centre/South		North		Difference (Centre/South-North)		
	Estimate	SE	Estimate	SE	Difference	SE	p.value
Debt Accumulation	-0.023	0.046	0.049	0.014	-0.072	0.048	0.134
Debt Repayment	0.020	0.035	-0.061	0.014	0.081	0.038	0.032
% Cost Overruns	0.185	0.063	0.006	0.021	0.179	0.066	0.007
Turnover Managers	0.133	0.051	0.075	0.032	0.058	0.060	0.335
% Graduate Managers	-0.144	0.030	-0.090	0.029	-0.054	0.042	0.196

Table F.12: RD estimates of the effect of electing a populist mayor estimated on two sub-samples of observations in the North and Centre-South of the country, as well as the difference in RD estimates. Same estimation of baseline analysis. To maximise sample size, no covariates were included. Northern regions are Veneto, Trentino-Alto Adige, Piemonte, Lombardia, Liguria, Friuli-Venezia Giulia, Emilia-Romagna. Centre and southern regions are the remaining 13 regions. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.

In Table F.13 below we show the difference in the effects estimated from a subset of the dataset used for the analysis with population below and above the median (8,616 inhabitants). Overall, we find smaller effects in larger municipalities for performance outcomes and no difference for turnover and the percentage of graduate managers.

Outcome	Above Median Population		Below Median Population		Difference (Above-Below)		
	Estimate	SE	Estimate	SE	Difference	SE	p.value
Debt Accumulation	-0.036	0.017	0.094	0.018	-0.130	0.025	0.000
Debt Repayment	0.026	0.015	-0.081	0.020	0.107	0.025	0.000
% Cost Overruns	0.018	0.027	0.101	0.041	-0.083	0.049	0.091
Turnover Managers	0.118	0.025	-0.131	0.363	0.249	0.364	0.494
% Graduate Managers	-0.108	0.021	-0.753	0.559	0.645	0.559	0.249

Table F.13: RD estimates of the effect of electing a populist mayor estimated on two sub-samples of observations in the North and Centre-South of the country, as well as the difference in RD estimates. Same estimation of baseline analysis. To maximise sample size, no covariates were included. Northern regions are Veneto, Trentino-Alto Adige, Piemonte, Lombardia, Liguria, Friuli-Venezia Giulia, Emilia-Romagna. Centre and southern regions are the remaining 13 regions. Period of analysis: cost overruns 2012-2020, debt accumulation and repayment 2008-2019, turnover and education of bureaucrats 2001-2019.